

Gauss Markov Proof

Gauss Markov: Given the assumptions of the classical linear regression model, the least squares estimators, in the class of unbiased linear estimators, have minimum variance, that is, they are BLUE.

Proof:

Linearity:

$$\hat{\beta}_1 = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2}$$

Let's multiply out the numerator:

$$\sum (X_i - \bar{X})(Y_i - \bar{Y}) = \sum (X_i - \bar{X})Y_i - \bar{Y}\sum (X_i - \bar{X}) = \sum (X_i - \bar{X})Y_i$$

$$\hat{\beta}_1 = \frac{\sum (X_i - \bar{X})Y_i}{\sum (X_i - \bar{X})^2} \quad \text{let } c_i = \frac{(X_i - \bar{X})}{\sum (X_i - \bar{X})^2}$$

$$\hat{\beta}_1 = \sum c_i Y_i$$

$\hat{\beta}_1$ is a weighted sum of the Y 's. As β_1 sums are linear, $\hat{\beta}_1$ is a linear estimator.

Unbiased

To be unbiased

$$E[\hat{\beta}_1] = \beta_1$$

$$\hat{\beta}_1 = \sum c_i Y_i$$

$$= \sum c_i (\beta_0 + \beta_1 X_i + \varepsilon_i)$$

$$= \beta_0 \sum c_i + \beta_1 \sum c_i X_i + \sum c_i \varepsilon_i$$

$$\sum c_i = \frac{\sum (X_i - \bar{X})}{\sum (X_i - \bar{X})^2} = 0$$

$$\frac{\sum (X_i - \bar{X}) \varepsilon_i}{\sum (X_i - \bar{X})^2} =$$

$$= \frac{\sum X_i \varepsilon_i}{\sum (X_i - \bar{X})^2} - \frac{\bar{X} \sum \varepsilon_i}{\sum (X_i - \bar{X})^2}$$

↑ by (5) ↑ by (2)

$$= \beta_1 \sum c_i X_i$$

$$= \beta_1 \sum c_i X_i + \beta_1 \sum c_i \bar{X} \leftarrow \text{add 0 b/c } \sum c_i = 0$$

$$\hat{\beta}_1 = \beta_1 \sum c_i (X_i - \bar{X})$$

$$E[\hat{\beta}_1] = E\left[\beta_1 \frac{\sum (X_i - \bar{X})^2}{\sum (X_i - \bar{X})^2}\right]$$

$$E[\hat{\beta}_1] = \beta_1$$

Efficiency

Proof by contradiction:

Imagine there are some weights (a_i) that are more efficient than the OLS weights c_i . Define

$$d_i \text{ as } d_i = a_i - c_i$$

Let our imaginary estimator be called \tilde{B}_1 .

$$\tilde{B}_1 = \sum a_i Y_i = \sum (c_i + d_i) Y_i$$

$$E[\tilde{B}_1] = E[\sum (c_i + d_i) Y_i] = E[\underbrace{\sum c_i Y_i}_{B_1}] + E[\sum d_i Y_i]$$

$$= B_1 + E[\sum d_i (B_0 + B_1 X_i + \varepsilon_i)] = B_1 + \sum d_i (B_0 + B_1 X_i)$$

$$E[\tilde{B}_1] = B_1 + B_0 \sum d_i + B_1 \sum d_i X_i$$

for \tilde{B}_1 to be unbiased $\sum d_i = 0$ & $\sum d_i X_i = 0$

$$\text{Var}(\tilde{B}_1) = \text{Var}[\sum a_i Y_i] = \text{Var}[\sum (c_i + d_i) Y_i] = \text{Var}[\sum c_i Y_i + \sum d_i Y_i]$$

Rule

$$\text{Let } Z = \sum b_i X_i$$

b_i are constants (weights)

$$\text{Then } \text{Var}(Z) = \sum b_i^2 \text{Var}(X_i) + 2 \sum_{i \neq j} b_i b_j \text{Cov}(X_i, X_j)$$

Proof: (put website)

$$Z = \sum b_i X_i$$

b_i are weights

$$\text{Var}(Z) = E \left[\sum b_i X_i - E \left[\sum b_i X_i \right] \right]^2$$

$$= E \left[\sum b_i X_i - \sum b_i E[X_i] \right]^2$$

$$= E \left[b_1 X_1 + b_2 X_2 + \dots + b_n X_n - b_1 E[X_1] - b_2 E[X_2] - b_3 E[X_3] - \dots - b_n E[X_n] \right]^2$$

$$= E \left[b_1 (X_1 - E[X_1]) + b_2 (X_2 - E[X_2]) + \dots + b_n (X_n - E[X_n]) \right]^2$$

$$= E \left[\sum_i b_i^2 (X_i - E[X_i])^2 + 2 \sum_{i \neq j} b_i b_j (X_i - E[X_i]) (X_j - E[X_j]) \right]$$

$$= \sum_i b_i^2 E (X_i - E[X_i])^2 + 2 \sum_{i \neq j} b_i b_j E \left[(X_i - E[X_i]) (X_j - E[X_j]) \right]$$

$$= \sum_i b_i^2 \text{Var}(X) + 2 \sum_{i \neq j} b_i b_j \text{Cov}(X_i, X_j)$$

apply rule $\text{Var}(z) = \sum b_i^2 \text{Var}(x) + 2 \sum \sum b_i b_j \text{Cov}(x_i, x_j)$
to $\text{Var}(\tilde{\beta}_1)$

$$\text{Var}(\tilde{\beta}_1) = \sum (c_i + d_i)^2 \text{Var}(Y_i) + 2 \sum_{i \neq j} (c_i + d_i)(c_j + d_j) \text{Cov}(Y_i, Y_j)$$

$$\text{Var}(Y_i) = \sigma^2 \text{ by assumption } \textcircled{3}$$

$$\text{Cov}(Y_i, Y_j) = 0 \text{ by assumption } \textcircled{4}$$

$$\text{Var}(\tilde{\beta}_1) = \sigma^2 \sum (c_i + d_i)^2$$

$$= \sigma^2 \sum c_i^2 + \sigma^2 \sum d_i^2 + 2\sigma^2 \sum c_i d_i$$

$$\frac{2\sigma^2}{\sum (x_i - \bar{x})^2} \left[\sum (x_i - \bar{x}) d_i \right]$$

$$= \frac{2\sigma^2}{\sum (x_i - \bar{x})^2} \left[\sum x_i d_i - \bar{x} \sum d_i \right]$$

$$\sum d_i = 0 \quad \sum x_i d_i = 0$$

$$\text{Var}(\tilde{\beta}_1) = \sum c_i^2 + \sigma^2 \sum d_i^2$$

This is minimized only when each $d_i = 0$ ✓